

Math 280C, Spring 2005

Exchangeability

Let $X = (X_1, X_2, \dots)$ be an exchangeable sequence of random variables. As discussed in class, there is no loss of generality (and some gain of convenience) in assuming that the sample space Ω is the sequence space $\mathbf{R}^{\mathbf{N}}$ (with generic element $\omega = (\omega_1, \omega_2, \dots)$) endowed with the product σ -field $\mathcal{F} := \mathcal{B}^{\mathbf{N}}$ (\mathcal{B} the Borel σ -field on \mathbf{R}), and that X is the sequence of coordinate random variables:

$$X_n(\omega) := \omega_n, \quad n = 1, 2, \dots$$

Notice that $\mathcal{F} = \sigma\{X_1, X_2, \dots\}$.

The tail σ -field \mathcal{T} is defined as

$$\mathcal{T} := \bigcap_n \mathcal{T}_n,$$

where

$$\mathcal{T}_n := \sigma\{X_{n+1}, X_{n+2}, \dots\}, \quad n = 1, 2, \dots$$

Let Σ_n denote the set of all permutations σ of \mathbf{N} with the property that $\sigma(k) = k$ for all $k > n$, and put $\Sigma = \bigcup_n \Sigma_n$, the set of all finite permutations of \mathbf{N} . Recall that the exchangeability of X means that

$$\sigma X := (X_{\sigma(1)}, X_{\sigma(2)}, \dots)$$

has the same distribution as X for each $\sigma \in \Sigma$. Put another way, we can view $\sigma \in \Sigma$ as inducing a mapping (also called σ) of Ω onto itself:

$$\sigma(\omega) := (\omega_{\sigma(1)}, \omega_{\sigma(2)}, \dots), \quad \omega \in \Omega,$$

and then exchangeability means that $\mathbf{P}[\sigma^{-1}(A)] = \mathbf{P}[A]$ for each $\sigma \in \Sigma$ and each $A \in \mathcal{F}$. (Here \mathbf{P} is the probability measure on (Ω, \mathcal{F}) governing X .) If $A \in \mathcal{F}$, then we say that A is n -symmetric provided $\sigma^{-1}A = A$ for every $\sigma \in \Sigma_n$. It is easy to check that the collection of such events forms a σ -field, which we denote \mathcal{E}_n . Evidently $\mathcal{E}_n \supset \mathcal{E}_{n+1}$; with this in mind we define the *exchangeable* σ -field \mathcal{E} as

$$\mathcal{E} := \bigcap_n \mathcal{E}_n.$$

Of course, \mathcal{E} is just the class of all events $A \in \mathcal{F}$ such that $\sigma^{-1}(A) = A$ for all $\sigma \in \Sigma$. Notice that

$$\mathcal{T}_n \subset \mathcal{E}_n, \quad \forall n \in \mathbf{N},$$

whence

$$\mathcal{T} \subset \mathcal{E}.$$

This inclusion is strict; for example, the event

$$A := \left\{ \omega : \sum_{k=1}^n X_k(\omega) \in [0, 1] \text{ for infinitely many } n \in \mathbf{N} \right\}.$$

lies in \mathcal{E} but not in \mathcal{T} . (Explain!) Nonetheless, for exchangeable sequences there is not much of a difference between \mathcal{T} and \mathcal{E} , as the following result of P.-A. Meyer shows. Its corollary, the Hewitt-Savage 0-1 law, originates in [1].

Theorem 1. *If $A \in \mathcal{E}$, then there exists $B \in \mathcal{T}$ such that $\mathbf{P}[A \Delta B] = 0$.*

Proof. Define $\mathcal{F}_n := \sigma\{X_1, X_2, \dots, X_n\}$, and $Z := 1_A$. By the martingale convergence theorems (forward and reverse):

$$\lim_n \mathbf{E}[Z|\mathcal{F}_n] = Z,$$

and

$$\lim_n \mathbf{E}[Z|\mathcal{T}_n] = \mathbf{E}[Z|\mathcal{T}].$$

in L^1 (and a.s.). Thus, given $\epsilon > 0$ there exists n_0 so large that

$$(1) \quad \|Z - \mathbf{E}[Z|\mathcal{F}_n]\|_1 < \epsilon$$

and

$$(2) \quad \|\mathbf{E}[Z|\mathcal{T}] - \mathbf{E}[Z|\mathcal{T}_n]\|_1 < \epsilon$$

provided $n \geq n_0$. For such an n let σ be the permutation that exchanges 1 and $n+1$, 2 and $n+2$, \dots , n and $2n$, and leaves the other integers alone. Because $A \in \mathcal{E}$, we have $Z \circ \sigma = Z$; because of this and the exchangeability of X , (1) implies

$$(3) \quad \|Z - \mathbf{E}[Z|\sigma\{X_{n+1}, \dots, X_{2n}\}]\|_1 < \epsilon.$$

But then, by the tower property and the L^1 -contraction property of conditional expectation,

$$(4) \quad \begin{aligned} \|\mathbf{E}[Z|\mathcal{T}_n] - \mathbf{E}[Z|\sigma\{X_{n+1}, \dots, X_{2n}\}]\|_1 &= \left\| \mathbf{E} \left[Z - \mathbf{E}[Z|\sigma\{X_{n+1}, \dots, X_{2n}\}] \middle| \mathcal{T}_n \right] \right\|_1 \\ &\leq \|Z - \mathbf{E}[Z|\sigma\{X_{n+1}, \dots, X_{2n}\}]\|_1 < \epsilon \end{aligned}$$

Combining (2), (3), and (4), we obtain

$$\|Z - \mathbf{E}[Z|\mathcal{T}]\|_1 < 3\epsilon.$$

Since $\epsilon > 0$ was arbitrary, it follows that

$$(5) \quad Z = \mathbf{E}[Z|\mathcal{T}] \text{ almost surely.}$$

The event $B := \{\omega : \mathbf{E}[Z|\mathcal{T}](\omega) = 1\}$ is \mathcal{T} -measurable, and (5) means that $\mathbf{P}[A\Delta B] = 0$. \square

Corollary. [Hewitt-Savage 0-1 law] *If the exchangeable sequence X is in fact iid, then \mathcal{E} is trivial: $\mathbf{P}[A] = 0$ or 1 for every $A \in \mathcal{E}$.*

Proof. By Kolmogorov's 0-1 law, the tail σ -field \mathcal{T} is trivial if the X_k s are independent. \square

The following result was mentioned in class. The general case, in which the random variables X_1, X_2, \dots are permitted to take values in a space more general than the real line, is due to E. Hewitt and L.J. Savage; see [1]. A splendid introduction to this result, and its applications, can be found in [2].

Theorem 2. [De Finetti's Theorem] *The sequence $X = (X_1, X_2, \dots)$ is conditionally iid, given \mathcal{E} . More precisely, if f_1, f_2, \dots, f_k are bounded and measurable, then*

$$\mathbf{E}[f_1(X_1)f_2(X_2)\cdots f_k(X_k)|\mathcal{E}] = \prod_{j=1}^k \mathbf{E}[f_j(X_1)|\mathcal{E}].$$

Proof. I shall write out the proof only for $k = 2$. The general case is not essentially different. By a symmetry argument used in class, we have (for $n \geq 2$)

$$(6) \quad \mathbf{E}[f_1(X_1)f_2(X_2)|\mathcal{E}_n] = \frac{1}{n(n-1)} \sum_{1 \leq i \neq j \leq n} f_1(X_i)f_2(X_j).$$

Now the right side of (6) differs from

$$\left\{ \frac{1}{n} \sum_{i=1}^n f_1(X_i) \right\} \left\{ \frac{1}{n} \sum_{j=1}^n f_2(X_j) \right\}$$

only in that the latter contains "diagonal" terms and the former does not. In fact, these two differ by no more than $2\|f\|_\infty^2/n$. It follows that, almost surely,

$$\begin{aligned} \mathbf{E}[f_1(X_1)f_2(X_2)|\mathcal{E}] &= \lim_n \mathbf{E}[f_1(X_1)f_2(X_2)|\mathcal{E}_n] = \lim_n \left\{ \frac{1}{n} \sum_{i=1}^n f_1(X_i) \right\} \left\{ \frac{1}{n} \sum_{j=1}^n f_2(X_j) \right\} \\ &= \lim_n \mathbf{E}[f_1(X_1)|\mathcal{E}_n] \cdot \mathbf{E}[f_2(X_1)|\mathcal{E}_n] \\ &= \mathbf{E}[f_1(X_1)|\mathcal{E}] \cdot \mathbf{E}[f_2(X_1)|\mathcal{E}], \end{aligned}$$

as claimed. \square

References

- [1] E. Hewitt and L.J. Savage: Symmetric measures on Cartesian products. *Trans. Amer. Math. Soc.* **80** (1955) 470–501.
- [2] J.F.C. Kingman: Uses of exchangeability. *Ann. Probability* **6** (1978) 183–197.