

## Math 280B, Winter 2008

### Conditioning and the Bivariate Normal Distribution

In what follows,  $X$  and  $Y$  are random variables defined on a probability space  $(\Omega, \mathcal{B}, P)$ , and  $\mathcal{G}$  is a sub- $\sigma$ -field of  $\mathcal{B}$ .

**1. Regular Conditional Distributions.** The conditional probability  $P[X \in B|\mathcal{G}]$  is defined to be the conditional expectation  $E[1_{\{X \in B\}}|\mathcal{G}] = E[1_B(X)|\mathcal{G}]$ , for  $B \in \mathcal{B}_{\mathbf{R}}$ . The function  $B \mapsto P[X \in B|\mathcal{G}]$  is “almost” a probability measure, in that  $P[X \in \mathbf{R}|\mathcal{G}] = 1$  almost surely and  $P[X \in \cup_{n=1}^{\infty} B_n|\mathcal{G}] = \sum_{n=1}^{\infty} P[X \in B_n|\mathcal{G}]$  almost surely for each sequence  $\{B_n\}$  of pairwise disjoint elements of  $\mathcal{B}_{\mathbf{R}}$ . The ambiguity present in these “almost surely” statements can be resolved because  $X$  is real-valued. (Similar considerations apply to a random variable with values in a measurable space that is measurably isomorphic to  $(\mathbf{R}, \mathcal{B}_{\mathbf{R}})$ .) This resolution permits a converse linkage between conditional probabilities and conditional expectations. The situation is summarized in the following result.

**Theorem.** *There is a function  $(\omega, B) \mapsto Q(\omega, B)$  from  $\Omega \times \mathcal{B}_{\mathbf{R}}$  to  $[0, 1]$  such that (i)  $\omega \mapsto Q(\omega, B)$  is  $\mathcal{G}$ -measurable for each  $B \in \mathcal{B}_{\mathbf{R}}$ , (ii)  $B \mapsto Q(\omega, B)$  is a probability measure on  $(\mathbf{R}, \mathcal{B}_{\mathbf{R}})$  for each  $\omega \in \Omega$ , and (iii) for each  $B \in \mathcal{B}_{\mathbf{R}}$ ,*

$$(1.1) \quad P[X \in B|\mathcal{G}](\omega) = Q(\omega, B) \quad \text{for } P\text{-a.e. } \omega \in \Omega.$$

Moreover, if  $\varphi : \mathbf{R} \rightarrow \mathbf{R}$  is Borel measurable and  $E|\varphi(X)| < \infty$ , then

$$(1.2) \quad E[\varphi(X)|\mathcal{G}](\omega) = \int_{\mathbf{R}} \varphi(x) Q(\omega, dx) \quad \text{for } P\text{-a.e. } \omega \in \Omega.$$

The function  $Q(\omega, B)$  is called a *regular conditional distribution for  $X$  given  $\mathcal{G}$* .

When  $\mathcal{G}$  is of the form  $\sigma(Y)$  there is a function  $(y, B) \mapsto F_y(B)$  from  $\mathbf{R} \times \mathcal{B}_{\mathbf{R}}$  to  $[0, 1]$  such that (i)  $y \mapsto F_y(B)$  is  $\mathcal{B}_{\mathbf{R}}$ -measurable for each  $B \in \mathcal{B}_{\mathbf{R}}$ , (ii)  $B \mapsto F_y(B)$  is a probability measure on  $(\mathbf{R}, \mathcal{B}_{\mathbf{R}})$  for each  $\omega \in \Omega$ , and (iii)  $Q(\omega, B) := F_{Y(\omega)}(B)$  is a regular conditional distribution for  $X$  given  $\sigma(Y)$ . It is natural to interpret  $F_y(B)$  as the conditional probability  $P[X \in B|Y = y]$ . A parallel interpretation of (1.2) is

$$(1.3) \quad E[\varphi(X)|Y = y] = \int_{\mathbf{R}} \varphi(x) F_y(dx).$$

**2. Basic Definition.** A pair  $(X, Y)$  of random variables, defined on some probability space  $(\Omega, \mathcal{F}, P)$ , is said to have a *bivariate normal* distribution (or to be jointly normally

distributed) provided the linear combination  $sX + tY$  is normally distributed for each pair  $(s, t) \in \mathbf{R}^2$ .

**3. Notation.** Let  $X$  and  $Y$  have a bivariate normal distribution. Taking  $s = 0$  and then  $t = 0$  in the Basic Definition, we see that the marginal distributions of  $X$  and  $Y$  are necessarily normal distributions. In particular,  $X$  and  $Y$  have moments of all orders. We use the following notation:

$$\mu := E[X], \sigma^2 := \text{Var}(X), \quad \nu := E[Y], \tau^2 := \text{Var}(Y),$$

and write

$$\rho = \rho(X, Y) := \text{Corr}(X, Y) = \text{Cov}(X, Y) / \sigma\tau$$

for the correlation of  $X$  and  $Y$ . Here  $\text{Cov}(X, Y) := E[(X - \mu)(Y - \nu)]$  is the covariance of  $X$  and  $Y$ .

**4. Characteristic Function.** In what follows  $(X, Y)$  will be a random vector with a bivariate normal distribution, and we shall use the notation of **3**. To avoid trivial cases we assume that  $\sigma > 0$  and  $\tau > 0$ . The (joint) characteristic function of  $X$  and  $Y$  is defined by

$$\phi_{X,Y}(s, t) := E[\exp(i(sX + tY))], \quad s, t \in \mathbf{R}.$$

In view of the Basic Definition,  $sX + tY \sim \mathcal{N}(s\mu + t\nu, s^2\sigma^2 + 2st\sigma\tau\rho + t^2\tau^2)$ , so

$$\phi_{X,Y}(s, t) = \exp \left[ i(s\mu + t\nu) - \frac{1}{2}(s^2\sigma^2 + 2st\sigma\tau\rho + t^2\tau^2) \right], \quad s, t \in \mathbf{R}.$$

**5. Independence.** Our goal is to compute explicitly the conditional distribution of  $X$  given  $Y$  in the bivariate normal case. We begin with a warm-up exercise: Compute the conditional expectation  $E[X|Y]$ . Our calculation is based on the following observation: By the Basic Definition, for any  $c \in \mathbf{R}$ , the pair  $(X - cY, Y)$  has a bivariate normal distribution. The covariance

$$\text{Cov}(X - cY, Y) = \text{Cov}(X, Y) - c\text{Cov}(Y, Y) = \sigma\tau\rho - c\tau^2$$

vanishes if and only if  $c = c^* := \sigma\rho/\tau$ . The random variables  $X - c^*Y$  and  $Y$  are then independent(!):

$$\begin{aligned} \phi_{X-c^*Y,Y}(s, t) &= \exp \left[ i(s\mu - sc^*\nu + t\nu) - \frac{1}{2}(s^2(\sigma^2 - 2c^*\sigma\tau\rho + [c^*]^2\tau^2) + t^2\tau^2) \right] \\ &= \exp \left[ is(\mu - sc^*\nu) - \frac{1}{2}s^2(\sigma^2 - 2c^*\sigma\tau\rho + [c^*]^2\tau^2) \right] \exp \left[ it\nu - \frac{1}{2}t^2\tau^2 \right] \\ &= \phi_{X-c^*Y}(s)\phi_Y(t). \end{aligned}$$

**6. Conditional Expectation.** In particular,

$$E[X|Y] = E[X - c^*Y|Y] + E[c^*Y|Y] = E[X - c^*Y] + c^*E[Y|Y] = \mu - c^*\nu + c^*Y,$$

where the second equality above follows from the independence of  $X - c^*Y$  and  $Y$ . We have shown that

$$E[X|Y] = \mu + \frac{\sigma\rho}{\tau}(Y - \nu), \quad \text{a.s.}$$

**7. Conditional Distribution, II.** I now claim that the conditional distribution  $F_y$  of  $X$  given  $Y = y$  (in the sense of **1** above) is the normal distribution with mean  $\mu + c^*(y - \nu)$  and variance  $(1 - \rho^2)\sigma^2$ . To see this let us use  $\Phi_y(s) = \exp(is(\mu + c^*(y - \nu)) - \frac{1}{2}s^2(1 - \rho^2)\sigma^2)$  to denote the associated characteristic function. Then, using the independence of  $X - c^*Y$  and  $Y$  for the third equality below:

$$\begin{aligned} E[\exp(isX)|Y] &= E[\exp(is(X - c^*Y)) \exp(isc^*Y)|Y] = E[\exp(is(X - c^*Y))|Y] \exp(isc^*Y) \\ &= E[\exp(is(X - c^*Y))] \exp(isc^*Y) \\ &= \exp \left[ is(\mu + c^*(Y - \nu)) - \frac{1}{2}s^2(\sigma^2 - 2c^*\sigma\tau\rho + [c^*]^2\tau^2) \right] \\ &= \exp \left[ is(\mu + c^*(Y - \nu)) - \frac{1}{2}s^2(\sigma^2 - 2\frac{\sigma\rho}{\tau}\sigma\tau\rho + \left(\frac{\sigma\rho}{\tau}\right)^2\tau^2) \right] \\ &= \exp \left[ is(\mu + c^*(Y - \nu)) - \frac{1}{2}s^2(\sigma^2 - \sigma^2\rho^2) \right] \\ &= \Phi_Y(s) \end{aligned}$$

It follows that for each  $s \in \mathbf{R}$ ,

$$E[\exp(isX)|Y](\omega) = \Phi_{Y(\omega)}(s) = \int_{\mathbf{R}} e^{isx} F_{Y(\omega)}(dx) \text{ for a.e. } \omega \in \Omega.$$

This confirms that  $\{F_y : y \in \mathbf{R}\}$  serves as a regular conditional distribution for  $X$  given  $Y$ .